



Bayesian Cox models for assessing *S. Typhimurium* virulence changes under an alternative antimicrobial treatment

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September 9, 2016

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Introduction

Background

- *Salmonella spp.* is one of the most relevant foodborne pathogens which causes about 155,000 salmonellosis-related deaths each year worldwide.
- *Salmonella enterica serovar Typhimurium* is one of the most usual serotypes related to salmonellosis outbreaks.
- Different kinds of **alternative preservation treatments** have been developed to reduce or eliminate this pathogen.
- The addition of **bioactive substances** from nature or agroindustrial by-products or the application of non-thermal treatments such as **High Hydrostatic Pressure** and **Pulsed Electric Fields** are the most usual antimicrobial alternatives.
- Most of the **alternative treatments** can generate **bacterial changes** or **adaptations** as a consequence of **repetitive** applications.

Experiment characterisation: aim

- **Evaluation** of the possible appearance of **bacterial adaptations** and **changes** caused by the **repetitive use** of an alternative preservation treatment in terms of virulence:
 - The not parasitic and transparent nematode, *C. elegans*, was used as the host system.
 - Virulence was measured in terms of the lifespan of the worm, egg laying and mobility.
 - We only focus here on lifespans.

Experiment characterisation: description

- **Three *S. Tyhimurium* populations:**
 - **Untreated** *S. Tyhimurium*.
 - *S. Tyhimurium* treated **one time**.
 - *S. Tyhimurium* treated **three times**.
- A population of 250 individuals of *C. elegans* **synchronised** by **age** was exposed to each of the three microbial populations.
- Nematodes were distributed in 25 plates of 10 worms.
- **Lifespan** of each nematode:
 - Measured in days.
 - Monitored each two or three days until the individual is dead.
 - Alive nematodes were transferred to a new plate at each monitoring time. This process can cause the accidental death of the nematode.

Modeling strategy

- *C. elegans* lifespan is modeled in each of the three populations of *S. Typhimurium* considered.
- **Survival regression model** with:
 - **Right censored** observations accidentally caused in the process of changing the worm from a plate to another one.
 - **Interval censored** observations because the event of interest, death of the worm, occurred between two consecutive monitoring times.
 - **Correlated observations** off the worms within the same plate.

Two types of models:

- **Cox Proportional Hazards (CPH)** models
- **Accelerated Failure Time (AFT)** models

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Two types of **models**:

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Statistical analysis: characterisation

Work in the context of **CPH models**:

- Allow to make inference about the parameters without assuming any distribution for the survival time (*semiparametric models*).
- Result in a natural estimate for the risk for failure associated with a vector of covariates.
- Widely accepted and utilized model for survival analysis.

Work in the context of the **right censored** data and ignoring interval censored by imputing the exact event time:

- Interval censored data analysis presents some implementation particularities in Bayesian software.
- **Interval censored observations** have been **imputed** into **non censored** through a random function defined in their respective interval time.

Work in the context of the **Bayesian Analysis**:

- Powerful and flexible methodology in relation to the modeling and quantification of uncertainties.
- Good in dealing with censored data.

Statistical analysis: Cox Proportional Hazard Model

- The CPH model assumes the following modeling for the hazard function

$$h(t | h_0, \mathbf{x}, \beta) = h_0(t) e^{\mathbf{x}'\beta}$$

where:

- $h_0(\cdot)$: **baseline hazard function**. It is a completely arbitrary hazard function that determines a baseline distribution with density $f_0(\cdot)$, cdf $F_0(\cdot)$, and survival function $S_0(\cdot)$.
- β is a vector of unknown parameters associated to covariates \mathbf{x} . The element $\mathbf{x}'\beta$ is known as the linear predictor.
- Much of the popularity of the CPH model is that Cox proposed methods for estimating the coefficients β without knowing $h_0(t)$.
- **Bayesian** inference **always** needs a model for $h_0(t)$.

- Our proposals for modeling $h_0(t)$ are based on:

1. **Weibull distribution**

- $h_0(t | \alpha, \lambda) = \lambda \alpha t^{\alpha-1}$.

2. **Piecewise function,**

- Time is divided into g prespecified intervals $I_k = (t_{k-1}, t_k]$ for $k = 1, 2, \dots, g$ where $0 = t_0 < t_1 < \dots < t_g < \infty$, being t_g the last survival or censored time.
- Baseline hazard is constant within intervals, $h_0(t_i) = \lambda_k$ for $t_i \in I_k$

Statistical analysis: Bayesian process

- Elicit a prior distribution for (h_0, β) but to do this we first need to model the baseline hazard function, $h_0(\cdot)$:
 1. Weibull.
 2. Piecewise:
 - 2.1 **Independent** λ 's
 - 2.2 λ 's correlated through a discrete-time **martingale process**.
 - 2.3 **Logarithmic transformation** of λ 's correlated through a discrete-time **martingale process**.
 - 2.4 λ 's correlated through a **Gamma process** prior.
- Construct the likelihood function of (β, h_0) for data $\mathcal{D} = \{(y_i, \delta_i, \mathbf{x}_i), i = 1, \dots, n\}$

$$L(h_0, \beta) = \prod_{i=1}^n L_i(h_0, \beta) = \prod_{i=1}^n [h_i(y_i | h_0, \beta)]^{\delta_i} [S_i(y_i | h_0, \beta)]$$

- Compute (approximate) the posterior distribution for (h_0, β)
- Compute posterior or predictive distributions for the relevant elements in the problem.

Statistical analysis: Model specification, I

- CPH model:

$$\blacksquare h_i(t \mid \mathbf{x}_i, \beta, h_0) = h_0(t) \exp\{\beta_{ST1} I_{ST1} + \beta_{ST3} I_{ST3}\}$$

- Baseline group: worms fed with untreated *S. Typhimurium*.
- $I_{ST1}(i) = 1$ if the worm i were fed with *S. Typhimurium* which survived one treatment application.
- $I_{ST3}(i) = 1$ if the worm i were fed with *S. Typhimurium* which survived three treatment applications.

Group	$h_i(t \mid \mathbf{x}_i, \beta, h_0)$
Control	$h_0(t)$
■ S.T treated one time	$h_0(t) \exp\{\beta_{ST1}\}$
S.T treated three times	$h_0(t) \exp\{\beta_{ST3}\}$

- Parameters of the model: β_{ST1} , β_{ST3} and the ones of the baseline hazard function.
 - $\pi(h_0, \beta) = \pi(h_0) \pi(\beta)$,
 - $\pi(\beta) = \prod_{j=1}^3 N(0, 0.01)$,
 - $\pi(h_0)$ in terms of a **Weibull** distribution or as a **piecewise function**.

Statistical analysis: Model specification, II

1. Weibull distribution

$$\alpha \sim \text{Ga}(1, 0.001), \lambda \sim \text{Ga}(0.001, 0.001)$$

2. Piecewise function

2.1 Independent λ 's

$$\lambda_k \sim \text{Ga}(0.01, 0.01), k = 1, \dots, g.$$

2.2 λ 's correlated through a discrete-time martingale process.

$$\lambda_1 \sim \text{Ga}(0.01, 0.01)$$

$$\lambda_k \mid \lambda_1, \dots, \lambda_{k-1} \sim \text{Ga}(0.01, 0.01/\lambda_{k-1}), k = 2, \dots, g.$$

2.3 Logarithmic transformation of λ 's correlated through a discrete-time martingale process.

$$\log(\lambda_k) = c_k$$

$$c_1 \sim \text{N}(0, \tau)$$

$$c_k \mid c_1, \dots, c_{k-1} \sim \text{N}(c_{k-1}, \tau), k = 2, \dots, g.$$

$$\tau \sim \text{Ga}(0.01, 0.01)$$

2.4 λ 's correlated through a Gamma process prior.

$$w_0 = 0.01; \eta_0 = 0.05$$

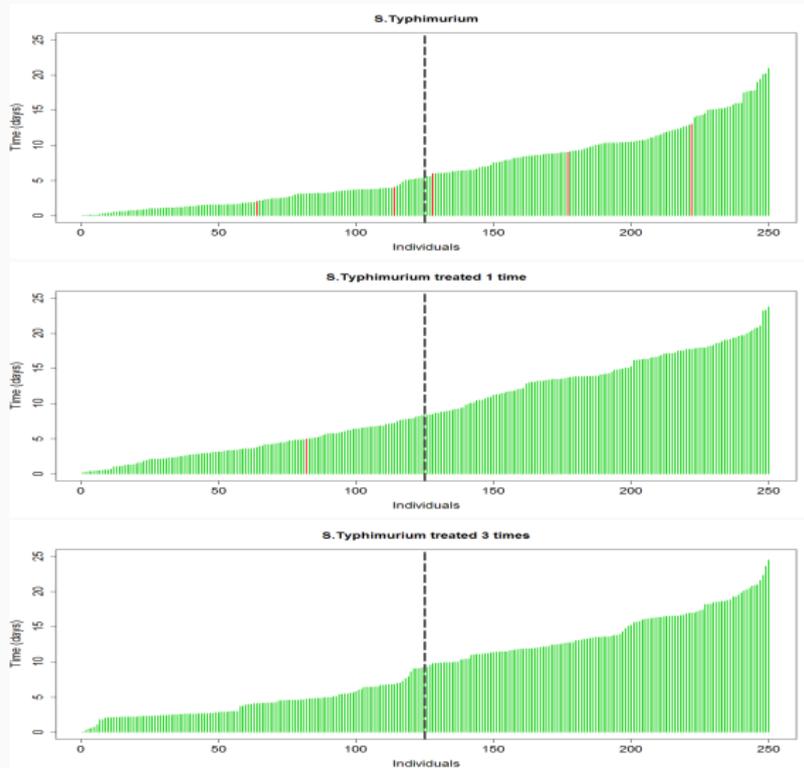
$$\lambda_1 \mid w_0 \sim \text{Ga}(w_0 10000, 10000)$$

$$\lambda_k \mid w_0, \eta_0 \sim \text{Ga}(w_0 \eta_0 ([t_{k+1} - t_k]), w_0 ([t_{k+1} - t_k])), k = 2, \dots, g.$$

Results

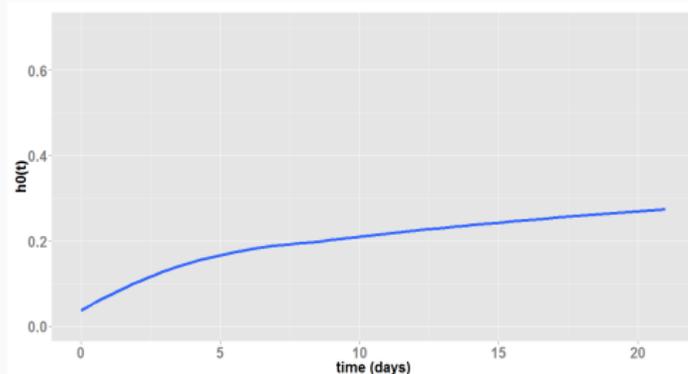
Data description

Survival time description

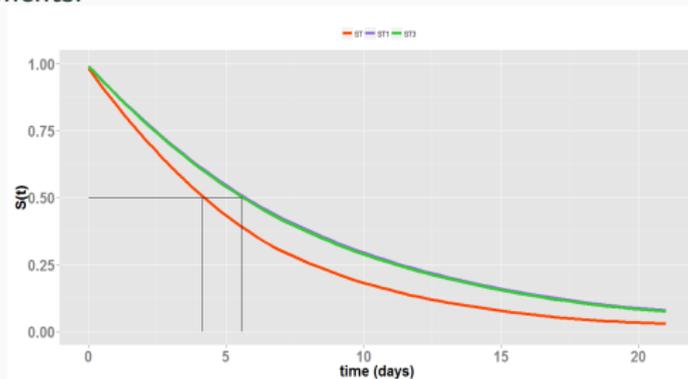


Weibull model

- Median of the posterior distribution for the hazard baseline risk function distribution $\pi(h_0(t) \mid \mathcal{D})$.

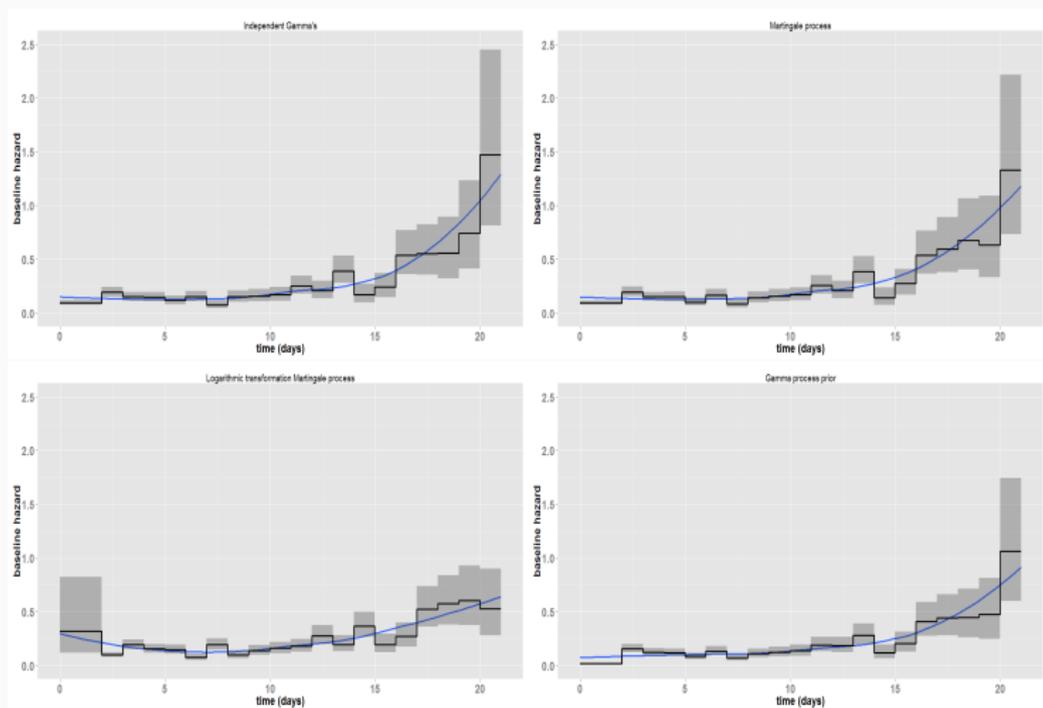


- Posterior median of the survival function distribution, $\pi(S(t \mid h_0, \beta) \mid \mathcal{D})$, for the three treatments.



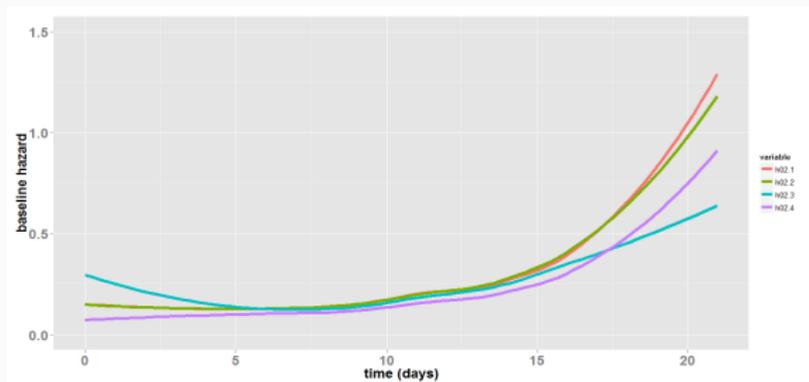
Piecewise function models, I.

- Median, smoothing median and 95% credible interval for the posterior distribution for the hazard baseline risk function $\pi(h_0(t) | \mathcal{D})$.



Piecewise function models, II.

- Smoothing median posterior distribution for the hazard baseline risk function $\pi(h_0(t) \mid \mathcal{D})$

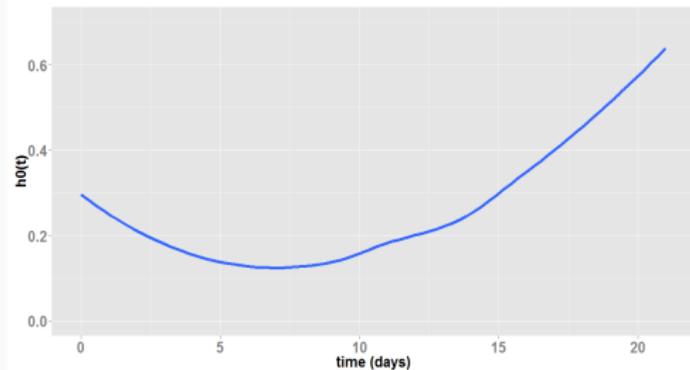


- Model selection:

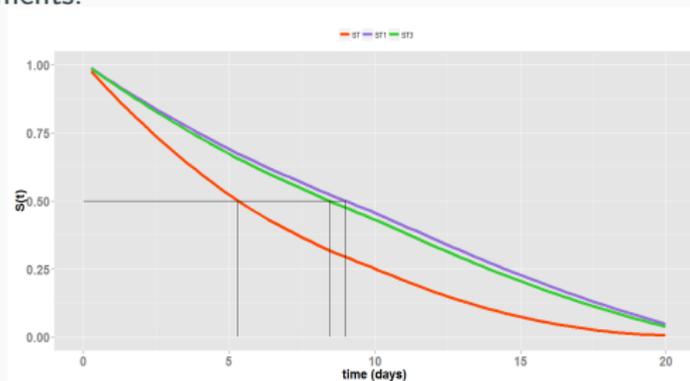
DIC	$h_{02.1}$	$h_{02.2}$	$h_{02.3}$	$h_{02.4}$
	5797.36	5746.11	5681.83	5924.94

Piecewise model

- Posterior median for the hazard baseline risk function distribution $\pi(h_0(t) | \mathcal{D})$

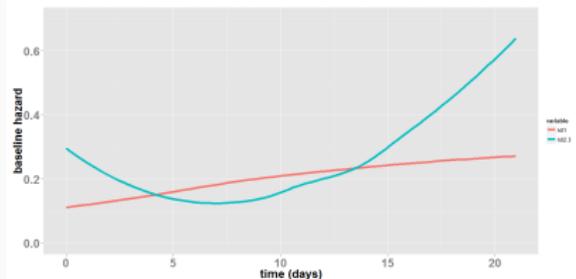


- Posterior median of the survival function distribution, $\pi(S(t | h_0, \beta) | \mathcal{D})$, for the three treatments.

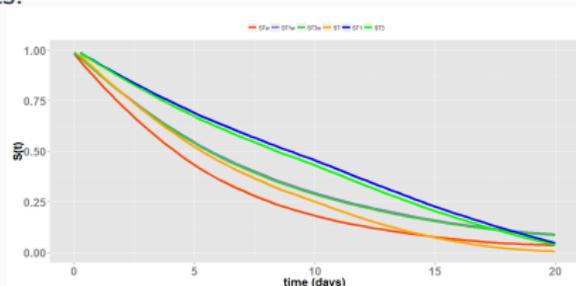


Model comparison

- Posterior median for the hazard baseline risk function distribution $\pi(h_0(t) | \mathcal{D})$



- Posterior median of the survival function distribution, $\pi(S(t | h_0, \beta) | \mathcal{D})$, for the three treatments.



- DIC:

Weibull	Piecewise
4553.4	5681.83

Conclusions and Future Research

Conclusions and Future Research

- **Conclusions:**

- Treated *S. Typhimurium* was less virulent than non-treated *S. Typhimurium*.
- Weak effect of frequency treatment.
- Piecewise models are heavily sensitive to prior specification for the baseline hazard model.
- Evident differences between Weibull and Piecewise model.
- Piecewise model based on Martingala logarithmic transformation seems to perform better.

- **Future Research:**

- Implement *semiparametric* models for interval censored data.
- Study different model selection criteria.
- Study sensitivity to number of intervals in the case of piecewise models.
- Include hierarchical structure in the models to correlate observations.
- And much more....

Questions?

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